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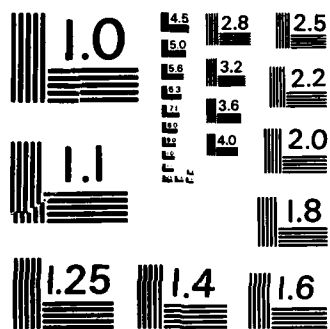
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REALIZED SIGNAL TO NOISE RATIO WITH AN ESTIMATED
DISCRIMINANT FUNCTION

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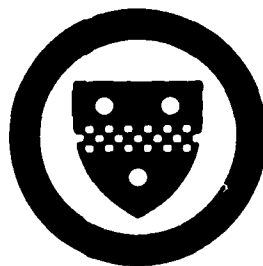
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ABSTRACT

an estimated discriminant function for signal detection is explained. *Keywords:* gender, variables, normal distribution, density functions, tables.

1. INTRODUCTION

Let X be a random variable having one of two possible p -variate normal distributions $N_p(0, \Sigma)$ and $N_p(\delta, \Sigma)$ when X is real (or $\tilde{N}_p(0, \Sigma)$ and $\tilde{N}_p(\delta, \Sigma)$ when X is complex). In the sequel, we follow the practice of giving the results for the real case first, and then for the complex case within brackets. For exact expressions of the p -variate normal density in the real and complex cases reference may be made to Khatri and Rao (1985). We consider the case where δ is known and Σ is unknown but an estimate $f^{-1}S$ of Σ

is available, where S has the Wishart distribution $W_p(f, \Sigma)$ (or $\tilde{W}_p(f, \Sigma)$) on f degrees of freedom. In such a case, the estimated linear discriminant function is proportional to $\delta'S^{-1}X$ (or $\delta^*S^{-1}X$). The discriminatory power or the probability of correct classification when the estimated discriminant function is used on future samples is a monotonic function of the signal to noise ratio

$$\rho(S, \Sigma) = \frac{(\delta'S^{-1}\delta)^2}{\delta'S^{-1}\Sigma S^{-1}\delta}, \quad (\text{or } \tilde{\rho}(S, \Sigma) = \frac{(\delta^*S^{-1}\delta)^2}{\delta^*S^{-1}\Sigma S^{-1}\delta}) \quad (1.1)$$

which involves the unknown Σ . By the Cauchy-Schwartz inequality, (1.1) is $\leq \delta'\Sigma^{-1}\delta$ (or $\delta^*\Sigma^{-1}\delta$), the signal to noise ratio when Σ is known so that there is some loss of information in using an estimated discriminant function. The problem is to make an inferential statement on the realized ratio (1.1) in terms of known quantities δ , S and f . This can be done in several ways as shown in the next section.

2. INFERENCE ON SIGNAL TO NOISE RATIO

The key result for this purpose is Theorem 1 in Khatri and Rao (1985) where it is shown that in the real case

$$B = \frac{(\delta'S^{-1}\delta)^2}{(\delta'\Sigma^{-1}\delta)(\delta'S^{-1}\Sigma S^{-1}\delta)} \quad \text{and} \quad G = \frac{\delta'\Sigma^{-1}\delta}{\delta'S^{-1}\delta} \quad (2.1)$$

are independently distributed with the p.d.f. (probability density function) of B as

$$\frac{\Gamma(\frac{f+1}{2})}{\Gamma(\frac{f-p+2}{2})\Gamma(\frac{p-1}{2})} b^{(f-p)/2} (1-b)^{(p-3)/2} \quad (2.2)$$

and that of G as

$$\frac{1}{2^{(f-p+1)/2}\Gamma(\frac{f-p+1}{2})} e^{-g/2} g^{(f-p+1)/2}. \quad (2.3)$$

In the complex case, defining \tilde{B} and \tilde{G} with δ' replaced by δ^* in (2.1), it is shown that \tilde{B} and \tilde{G} are independently distributed, with the p.d.f. of \tilde{B} as

$$\frac{\Gamma(f+1)}{\Gamma(f-p+2)\Gamma(p-1)} \tilde{b}^{(f-p+1)} (1-\tilde{b})^{p-2} \quad (2.4)$$

and that of \tilde{G} as

$$\frac{1}{\Gamma(f-p+1)} e^{-g} g^{f-p}. \quad (2.5)$$

The distribution (2.4) was earlier obtained by Reed, Mallet and Brennan (1974). The distributions (2.3)-(2.5) are independent of the unknown parameters which enables inferences to be drawn on (1.1) through the pivotal statistics B and G (or \tilde{B} and \tilde{G}).

1) Using the expressions for the moments of the beta distribution (Rao, 1973, p. 168)

$$E[B(\text{or } \tilde{B})] = \frac{f-p+2}{f+1}$$

which gives

$$E[\rho(S, \Sigma)(\text{or } \tilde{\rho}(S, \Sigma))] = \frac{f-p+2}{f+1} \{\delta' \Sigma^{-1} \delta (\text{or } \delta^* \Sigma^{-1} \delta)\}. \quad (2.6)$$

If the average efficiency is to be maintained at about half the optimal efficiency, then from (2.6) we have

$$\frac{f-p+2}{f+1} = \frac{1}{2} \text{ or } f = 2p \quad (2.7)$$

for both the real and complex cases. The result for the complex case is mentioned in Reed, Mallet and Brennan (1974). Similarly we can equate the ratio in (2.7) to any desired ratio other than $(\frac{1}{2})$ and obtain an expression for the degrees of freedom f needed for the estimation of Σ .

2) Perhaps a more satisfactory way of using the distributions (2.2) and (2.4) is as follows. Let b_α (or \tilde{b}_α) be the lower $\alpha\%$ point of

the distribution (2.2) (or (2.4)). Then we can make the confidence statement that

$$\rho(S, \Sigma) \geq b_{\alpha} \delta' \Sigma^{-1} \delta \quad (\text{or } \tilde{\rho}(S, \Sigma) \geq \tilde{b}_{\alpha} \delta' \Sigma^{-1} \delta) \quad (2.8)$$

with a confidence coefficient of $(1-\alpha)\%$.

3) The results (2.7) and (2.8) still involve the unknown quantity Σ . We raise the question whether the actual magnitude of $\rho(S, \Sigma)$ for given S can be assessed through known quantities. Using the joint distribution of B and G as in (2.2) and (2.3), or \tilde{B} and \tilde{G} as in (2.4) and (2.5), it is shown in Khatri and Rao (1985) that the estimate

$$\hat{\rho}(S, \Sigma) = \frac{(f-p+2)(f-p-3)}{f(f+1)} D_p^2 \quad (\text{or } \hat{\tilde{\rho}}(S, \Sigma) = \frac{(f-p+2)(f-p-1)}{f(f+1)} D_p^2)$$

where $D_p^2 = f \delta' S^{-1} \delta$ (or $f \delta' \Sigma^{-1} \delta$) has the property $\text{Min}_g E[\rho(S, \Sigma) - g D_p^2]^2$

$$= E[\rho(S, \Sigma) - \hat{\rho}(S, \Sigma)]^2 \quad (\text{or } = E[\tilde{\rho}(S, \Sigma) - \hat{\tilde{\rho}}(S, \Sigma)]^2).$$

so that $\hat{\rho}(S, \Sigma)$ (or $\hat{\tilde{\rho}}(S, \Sigma)$) is a predictor of $\rho(S, \Sigma)$ (or $\tilde{\rho}(S, \Sigma)$)

4) Khatri and Rao (1985) also obtained an exact confidence bound for $\rho(S, \Sigma)$ for the computation of which we provide extensive tables in this paper.

We define the random variable

$$Z = \frac{1}{2} BG = \frac{f}{2} \frac{\rho(S, \Sigma)}{D_p^2} \quad (\text{or } \tilde{Z} = f \frac{\tilde{\rho}(S, \Sigma)}{D_p^2})$$

which has the confluent hypergeometric distribution with the probability density function

$$\frac{e^{-z} z^{m-1}}{\Gamma(m)} \frac{\Gamma(a+b)}{\Gamma(a)} \Psi(b, m-a+1; z).$$

where $m = (f-p+1)/2$, $a = (f-p+2)/2$, $b = (p-1)/2$, $m-a+1 = 1/2$

(or $m = f-p+1$, $a = f-p+2$, $b = p-1$, $m-a+1 = 0$), and

$$\psi(b, c; z) = \frac{1}{\Gamma(b)} \int_0^\infty t^{b-1} (1+t)^{c-b-1} \exp(-zt) dt$$

as described in Khatri and Rao (1985, Theorem 3 and Remark 5). If z_α (or \tilde{z}_α) is the lower $\alpha\%$ point of this distribution, then

$$\rho(S, \Sigma) \geq \frac{2z_\alpha}{f} D_p^2, \text{ (or } \tilde{\rho}(S, \Sigma) \geq \frac{\tilde{z}_\alpha}{f} D_p^2)$$

provides the lower confidence bound to the signal to noise ratio $\rho(S, \Sigma)$ (or $\tilde{\rho}(S, \Sigma)$) with a confidence coefficient of $(1-\alpha)\%$. Tables 1-5 given in the next section give the values of $2z_\alpha$ (or \tilde{z}_α) for various combinations of p and f , and $\alpha = 0.05, 0.25, 0.50, 0.75$ and 0.95 .

3. PERCENTAGE POINTS AND SOME APPROXIMATIONS

Tables 1-5 give the lower $\alpha\%$ (for $\alpha\% = 5, 25, 50, 75$ and 95) points of the distributions of $2z$ and \tilde{z} for different-values of p and f . Actually z_α is obtained as a solution to the equation

$$\alpha = \int_0^1 \frac{\Gamma(\frac{f+1}{2})}{\Gamma(\frac{p-1}{2})\Gamma(\frac{f-p+2}{2})} y^{(f-p)/2} (1-y)^{(p-3)/2} dy \int_0^{z_\alpha/y} \frac{1}{\Gamma(\frac{f-p+1}{2})} e^{-x} x^{(f-p-1)/2} dx$$

and \tilde{z}_α to the equation

$$\alpha = \int_0^1 \frac{\Gamma(f+1)}{\Gamma(p-1)\Gamma(f-p+2)} y^{f-p+1} (1-y)^{p-2} dy \int_0^{\tilde{z}_\alpha/y} \frac{1}{\Gamma(f-p+1)} e^{-x} x^{f-p} dx.$$

To obtain $(1-\alpha)\%$ lower confidence bound, we multiply the Mahalanobis distance $D_p^2 = f\delta'S^{-1}\delta$ (or $f\tilde{\delta}*\tilde{S}^{-1}\tilde{\delta}$) by $2z_\alpha/f$ (or \tilde{z}_α/f).

We give several approximations to the distributions of z (or \tilde{z}) from which fairly approximate values of z_α (or \tilde{z}_α) can be easily obtained if p/f is not large.

(i) Gamma approximation

(a) We consider the statistic

$$g(p,f)z \text{ (or } \tilde{g}(p,f)\tilde{z}) \quad (3.1)$$

and approximate its distribution by a gamma distribution $G(1,v)$ or $G(1,\tilde{v})$.

To determine $g(p,f)$ and v we equate the first two moments of $g(p,f)z$ with those of $G(1,v)$. The equations are

$$E[g(p,f)z] = g(p,f)E(z) = v$$

$$V[g(p,f)z] = [g(p,f)]^2 V(z) = v$$

which give

$$g(p,f) = \frac{E(z)}{V(z)} = \frac{(f+1)(f+3)}{(f+1)(f+3)-p(p-1)}$$

$$v = \frac{[E(z)]^2}{V(z)} = \frac{f-p+1}{2} \frac{(f+3)(f-p+2)}{(f+1)(f+3)-p(p-1)}$$

Similarly

$$\tilde{g}(p,f) = \frac{E(\tilde{z})}{V(\tilde{z})} = \frac{(f+1)(f+2)}{(f+1)(f+2)-p(p-1)}$$

$$\tilde{v} = \frac{[E(\tilde{z})]^2}{V(\tilde{z})} = \frac{(f+2)(f-p+1)(f-p+2)}{(f+1)(f+2)-p(p-1)}$$

(b) We consider the statistic

$$g(p,f)z^c \text{ (or } \tilde{g}(p,f)\tilde{z}^c) \quad (3.2)$$

and approximate its distribution by a gamma distribution $G(1,v)$ (or $G(1,\tilde{v})$). To determine $g(p,f)$, c and v , we equate the first three moments of $g(p,f)z^c$ with those of $G(1,v)$. Taking $m = (f-p+1)/2$, $a = (f-p+2)/2$ and $b = (p-1)/2$, the equations are

$$v = g(p, f) \Gamma(m+c) \Gamma(a+c) \Gamma(a+b) / \{\Gamma(m) \Gamma(a) \Gamma(a+b+c)\},$$

$$v+1 = g(p, f) \Gamma(m+2c) \Gamma(a+2c) \Gamma(a+b+c) / \{\Gamma(m+c) \Gamma(a+c) \Gamma(a+b+2c)\}$$

$$v+2 = g(p, f) \Gamma(m+3c) \Gamma(a+3c) \Gamma(a+b+2c) / \{\Gamma(m+2c) \Gamma(a+2c) \Gamma(a+b+3c)\}.$$

These equations give

$$\begin{aligned} \frac{\Gamma(m+3c) \Gamma(a+3c) \Gamma(a+b+2c)}{\Gamma(m+2c) \Gamma(a+2c) \Gamma(a+b+3c)} + \frac{\Gamma(m+c) \Gamma(a+c) \Gamma(a+b)}{\Gamma(m) \Gamma(a) \Gamma(a+b+c)} \\ = 2 \frac{\Gamma(m+2c) \Gamma(a+2c) \Gamma(a+b+2c)}{\Gamma(m+c) \Gamma(a+c) \Gamma(a+b+2c)} \end{aligned}$$

from which a solution for c is computed using an appropriate computer program.

Similarly, for the complex situation, we take $m = f-p+1$, $a = f-p+2$ and $b = p-1$, and determine the values of \tilde{c} , $\tilde{g}(p, f)$ and \tilde{v} as above. The values of these constants for the real and complex cases are given in Table 6 for some values of p and f .

We find by comparing with actual values that the approximation given by (3.2) is more accurate than that given by (3.1), and the approximation is fairly accurate even for small values of f and p .

(ii) Normal approximation

(a) Using the Wilson-Hilferty's approximation as modified by Konishi (1981), we have

$$(9v)^{1/2} \left\{ \left(\frac{g(p, f)z}{2v} \right)^{1/3} - 1 + \frac{1}{9v} \right\} \sim N(0, 1)$$

$$\left[\text{or } (9\tilde{v})^{1/2} \left\{ \left(\frac{\tilde{g}(p, f)\tilde{z}}{2\tilde{v}} \right)^{1/3} - 1 + \frac{1}{9\tilde{v}} \right\} \sim N(0, 1) \right].$$

where $g(p, f)$ and v (or $\tilde{g}(p, f)$ and \tilde{v}) are defined in (3.1).

(b) As in (a), we can take

$$(9v)^{1/2} \left\{ \left(\frac{g(p, f)z^c}{2v} \right)^{1/3} - 1 + \frac{1}{9v} \right\} \sim N(0, 1)$$

$$\left[\text{or } (9\tilde{v})^{1/2} \left\{ \left(\frac{\tilde{g}(p, f)\tilde{z}^c}{2\tilde{v}} \right)^{1/3} - 1 + \frac{1}{9\tilde{v}} \right\} \sim N(0, 1) \right]$$

where $g(p, f)$, v and c (or $\tilde{g}(p, f)$, \tilde{v} and \tilde{c}) are defined in (3.2). The normal approximation is fairly accurate even for small values of f and p .

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TABLE I. Values of $2z_\alpha$ and z_α for $\alpha = 0.05$ and various combinations of p (number of variables) and f (degrees of freedom). ($2z_\alpha$ is the upper value and z_α is the lower value)

$p \backslash f$	5	10	15	20	25	30	35	40	45	50	60	70	80	90	100
2	0.512	2.876	5.952	9.232	13.150	16.976	20.900	24.900	29.026	33.100	41.524	50.048	58.678	67.394	76.184
3	1.050	4.098	7.813	11.725	15.775	19.950	24.206	28.507	32.854	37.239	46.099	55.056	64.088	73.181	82.325
4	0.172	2.026	4.900	8.214	11.788	15.550	19.450	23.426	27.450	31.550	39.896	48.386	56.992	65.688	74.458
5	0.456	3.064	6.488	10.288	14.275	18.413	22.604	26.872	31.193	35.555	44.382	53.312	62.324	71.401	80.531
6	0.030	1.338	3.912	7.052	10.500	14.176	18.026	21.950	25.950	30.026	38.304	46.758	55.332	64.004	72.756
7	0.131	2.164	5.325	8.975	12.875	16.925	21.063	25.291	29.579	33.915	42.699	51.599	60.587	69.644	78.758
8	---	0.832	3.076	6.028	9.326	12.900	16.650	20.500	24.500	28.500	36.748	45.160	53.700	62.346	71.074
9	---	1.450	4.288	7.750	11.525	15.475	19.581	23.762	28.011	32.316	41.053	49.917	58.876	67.911	77.007
10	---	0.460	2.350	5.078	8.226	11.700	15.350	19.150	23.050	27.050	35.230	43.592	52.096	60.710	69.414
11	---	0.894	3.375	6.625	10.275	14.150	18.159	22.285	26.491	30.760	39.441	48.264	57.192	66.202	75.277
12	---	0.216	1.726	4.214	7.176	10.550	14.100	17.850	21.700	25.650	33.746	42.054	50.518	59.100	67.776
13	---	0.488	2.575	5.613	9.088	12.850	16.796	20.860	25.017	29.245	37.864	46.641	55.534	64.516	73.568
14	---	0.076	1.226	3.450	6.226	9.450	12.950	16.576	20.376	24.300	32.706	40.548	48.966	57.512	66.160
15	---	0.217	1.906	4.675	7.994	11.650	15.491	19.487	23.589	27.773	36.322	45.049	53.902	62.853	71.879
16	---	0.014	0.826	2.776	5.350	8.450	11.826	15.376	19.100	22.950	30.884	39.072	47.442	55.948	64.564
17	---	0.064	1.338	3.838	6.975	10.488	14.244	18.164	22.206	26.341	34.815	43.486	52.297	61.213	70.212
18	---	---	0.512	2.176	4.550	7.476	10.750	14.226	17.900	21.650	29.506	37.626	45.944	54.408	62.990
19	---	---	0.884	3.094	6.038	9.413	13.054	16.893	20.869	24.951	33.343	41.953	50.717	59.597	68.566
20	---	---	0.288	1.676	3.812	6.588	9.726	13.100	16.700	20.426	28.162	36.210	44.472	52.892	61.436
21	---	---	0.531	2.438	5.175	8.388	11.921	15.671	19.577	23.601	31.904	40.449	49.164	58.004	66.940
22	---	---	0.136	1.238	3.150	5.776	8.776	12.050	15.550	19.226	26.854	34.824	43.026	51.398	59.906
23	---	---	0.278	1.863	4.375	7.425	10.844	14.500	18.330	22.292	30.500	38.976	47.636	56.433	65.335
24	---	---	0.048	0.882	2.576	5.012	7.876	11.050	14.450	18.050	25.580	33.468	41.606	49.930	58.394
25	---	---	0.113	1.375	3.650	6.538	9.822	13.378	17.128	21.024	29.130	37.531	46.134	54.886	63.752
26	---	---	0.008	0.594	2.050	4.300	7.026	10.100	13.426	16.900	24.340	32.142	40.214	48.482	56.904
27	---	---	0.026	0.969	3.006	5.713	8.856	12.305	15.969	19.796	27.794	36.116	44.659	53.362	62.188
28	---	---	---	0.376	1.594	3.650	6.250	9.188	12.426	15.850	23.134	30.844	38.846	47.060	55.436
29	---	---	---	0.644	2.425	4.938	7.944	11.280	14.855	18.607	26.492	34.730	43.208	51.861	60.646
30	---	---	---	---	0.228	1.288	3.050	5.350	8.026	11.026	17.606	24.796	32.396	40.290	48.404
31	---	---	---	---	0.503	1.988	4.179	6.874	9.931	13.257	20.483	28.234	36.340	44.697	53.242
32	---	---	---	---	---	0.184	1.076	2.626	4.676	7.126	12.896	19.464	26.580	34.088	41.888
33	---	---	---	---	---	0.413	1.681	3.621	6.059	8.870	15.295	22.454	30.103	38.098	46.349
34	---	---	---	---	---	---	0.152	0.926	2.300	4.150	8.976	14.830	21.386	28.446	35.880
35	---	---	---	---	---	---	0.348	1.458	3.195	5.416	10.911	17.375	24.488	32.057	39.961
36	---	---	---	---	---	---	---	0.130	0.806	2.038	5.820	10.872	16.800	23.354	30.372
37	---	---	---	---	---	---	---	0.302	1.288	2.858	7.309	12.985	19.486	26.567	34.073
38	---	---	---	---	---	---	---	---	0.114	0.718	3.396	7.576	12.808	18.800	25.356
39	---	---	---	---	---	---	---	---	0.266	1.153	4.468	9.271	15.088	21.622	28.680
40	---	---	---	---	---	---	---	---	---	0.100	1.666	4.916	9.398	14.776	20.826
41	---	---	---	---	---	---	---	---	---	0.239	2.360	6.215	11.282	17.213	23.777
42	---	---	---	---	---	---	---	---	---	---	0.582	2.870	6.552	11.272	16.772
43	---	---	---	---	---	---	---	---	---	---	---	---	8.059	13.333	19.357

TABLE VI. Values of c , g , f (real case) and \tilde{c} , \tilde{g} , \tilde{f} (complex case) for the chi-square approximation (3.2)

$f \setminus p$	2	3	4	5	6	7	8	9	10	11	12	13	14	15	20
c	0.952	0.878	0.790	---	---	---	---	---	---	---	---	---	---	---	---
g	1.194	1.549	2.078	---	---	---	---	---	---	---	---	---	---	---	---
f	1.920	1.488	1.089	---	---	---	---	---	---	---	---	---	---	---	---
\tilde{c}	0.943	0.862	0.772	---	---	---	---	---	---	---	---	---	---	---	---
\tilde{g}	1.272	1.775	2.517	---	---	---	---	---	---	---	---	---	---	---	---
\tilde{f}	3.934	3.150	2.381	---	---	---	---	---	---	---	---	---	---	---	---
c	0.980	0.948	0.907	0.861	0.811	0.759	0.706	0.653	---	---	---	---	---	---	---
g	1.087	1.244	1.468	1.765	2.140	2.597	3.133	3.732	---	---	---	---	---	---	---
f	4.316	3.801	3.379	3.005	2.644	2.270	1.854	1.366	---	---	---	---	---	---	---
\tilde{c}	0.978	0.942	0.898	0.850	0.798	0.746	0.695	0.648	---	---	---	---	---	---	---
\tilde{g}	1.112	1.321	1.622	2.030	2.554	3.201	3.958	4.773	---	---	---	---	---	---	---
\tilde{f}	8.682	7.726	6.945	6.250	5.567	4.838	3.996	2.969	---	---	---	---	---	---	---
c	0.989	0.971	0.947	0.919	0.887	0.854	0.819	0.783	0.747	0.711	0.674	0.638	0.602	---	---
g	1.050	1.142	1.272	1.442	1.653	1.907	2.208	2.559	2.961	3.413	3.908	4.438	4.978	---	---
f	6.751	6.162	5.675	5.257	4.881	4.529	4.185	3.835	3.468	3.069	2.623	2.115	1.525	---	---
\tilde{c}	0.990	0.968	0.940	0.914	0.878	0.846	0.810	0.772	0.737	0.701	0.666	0.633	0.603	---	---
\tilde{g}	1.053	1.181	1.359	1.555	1.856	2.184	2.597	3.104	3.658	4.311	5.013	5.763	6.477	---	---
\tilde{f}	13.479	12.412	11.537	10.659	10.025	9.319	8.665	8.023	7.276	6.490	5.576	4.527	3.276	---	---
c	0.993	0.982	0.965	0.946	0.924	0.901	0.876	0.850	0.823	0.796	0.768	0.741	0.713	0.685	---
g	1.036	1.093	1.182	1.293	1.433	1.595	1.790	2.013	2.271	2.559	2.883	3.246	3.643	4.074	---
f	9.219	8.555	8.021	7.547	7.135	6.749	6.397	6.055	5.724	5.382	5.034	4.670	4.280	3.856	---
\tilde{c}	1.000	0.984	0.962	0.941	0.924	0.894	0.869	0.842	0.817	0.787	0.760	0.731	0.706	0.678	---
\tilde{g}	1.002	1.096	1.229	1.376	1.512	1.765	2.015	2.313	2.641	3.066	3.501	4.034	4.553	5.198	---
\tilde{f}	18.154	17.048	16.160	15.290	14.320	13.768	13.077	12.427	11.736	11.146	10.430	9.750	8.899	8.083	---
c	0.996	0.987	0.975	0.963	0.946	0.928	0.910	0.890	0.868	0.848	0.825	0.804	0.781	0.759	0.647
g	1.021	1.066	1.132	1.204	1.307	1.426	1.562	1.720	1.902	2.099	2.326	2.575	2.850	3.155	5.066
f	11.662	10.980	10.401	9.849	9.408	9.002	8.615	8.260	7.926	7.584	7.260	6.926	6.589	6.242	4.153
\tilde{c}	0.988	0.998	0.990	0.958	0.931	0.929	0.911	0.881	0.859	0.846	0.825	0.796	0.773	0.752	0.642
\tilde{g}	1.066	1.018	1.069	1.269	1.466	1.497	1.654	1.931	2.173	2.348	2.637	3.053	3.433	3.834	6.582
\tilde{f}	23.700	21.488	20.197	19.936	19.488	18.025	17.235	16.891	16.259	15.304	14.657	14.241	13.581	12.853	8.683

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20. ABSTRACT (Continue on reverse side if necessary and identify by block number) Percentage points of a new distribution involving a confluent-hypergeometric distribution obtained by Khatri and Rao (1985) are tabulated. The use of the tabulated values in obtaining a lower confidence bound for the realized signal to noise ratio based on an estimated discriminant function for signal detection is explained.		

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